

Lecture 17: Stochastic Bandits & UCB

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Disclaimer: These notes have not been subjected to the usual scrutiny reserved for formal publications.

17.1 Notation

- $\mathbb{E}[\cdot]$: Expectation
- $\Pr[B]$: Probability of event B
- $\mathbb{1}\{B\}$: Indicator function for the event B

17.2 Stochastic Bandits

Setup

Consider the situation where there exists n probability distributions D_1, \dots, D_n , and let μ_i be the mean of D_i , for all $i = 1, \dots, n$. In the multi-arm bandit setting with n arms, in round t , arm i pays $X_i^t \stackrel{\text{iid}}{\sim} D_i$ (note: iid in rounds). Typically, it is assumed that a distribution D_i is 1-sub-Gaussian. The assumption is to ensure that the empirical and the true mean of the distributions are within certain bound with some probability.

Now, an algorithm for the multi-arm bandit problem selects $i_t \in \{1, \dots, n\}$ at round t , while receiving a reward $X_{i_t}^t$, and can only observe that reward ($X_{i_t}^t$) alone. For the sake of analysis, without loss of generality, assume

$$\mu_1 = \max_{i \in \{1, \dots, n\}} \mu_i, \quad (17.1)$$

and let $\Delta_i = \mu_1 - \mu_i$, $i = 2, \dots, n$. Finally, assume that the value of $\Delta_* = \min_{i \in \{2, \dots, n\}} \Delta_i$ is known.

A simple algorithm to minimize expected regret over T rounds, which is defined as

$$\mathbb{E}[\text{regret}_T] = \sum_{i=2}^n \Delta_i \mathbb{E}[N_i^{T+1}], \quad (17.2)$$

where $N_i^t = \sum_{s=1}^{t-1} \mathbb{1}\{i_s = i\}$, is discussed below. Note that N_i^t denotes the number of times arm i is pulled up till the t^{th} round.

Simple Algorithm

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if  $t \in [(i-1)k + 1, k]$  then
   $i_t = i$ 
end if
if  $t > nk$  then
   $i_t = \arg \max_{i \in \{1, \dots, n\}} \hat{\mu}_i$ 
end if

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Here,

$$k = \left\lceil \frac{4 \log(nT)}{\Delta_*^2} \right\rceil, \quad (17.3)$$

and

$$\hat{\mu}_i = \frac{1}{k} \sum_{t=(i-1)k+1}^{ik} X_i^t.$$

Theorem 17.1 *The bound on the expected regret for the simple algorithm over T rounds is given by*

$$\mathbb{E}[\text{regret}_T] \leq \sum_{i=2}^n \left[\frac{4\Delta_i \log(nT)}{\Delta_*^2} + O(\Delta_i) \right]. \quad (17.4)$$

Proof: We prove the above statement by first finding a bound to $\mathbb{E}[N_i^{T+1}]$, for $i \neq 1$. Note that

$$\mathbb{E}[N_i^{T+1}] = \mathbb{E}[N_i^{T+1} \mathbb{1}\{B\} + N_i^{T+1} \mathbb{1}\{\bar{B}\}], \quad (17.5)$$

where $B = \{\hat{\mu}_1 > \hat{\mu}_i, \forall i = 2, \dots, n\}$, and \bar{B} represents the nonoccurrence of event B . If event B takes places, then arm 1 will be chosen after kn rounds, then all the other arms would have been chosen atmost k times. Subsequently $\mathbb{E}[N_i^{T+1} \mathbb{1}\{B\}] \leq k$. Clearly, the upper bound for the latter term can be $\mathbb{E}[N_i^{T+1} \mathbb{1}\{\bar{B}\}] \leq T \Pr[\bar{B}]$.

We now obtain an upper bound for $\Pr[\bar{B}]$ by first realizing

$$\begin{aligned} \Pr[\bar{B}] &= \Pr[\exists i \in \{2, \dots, n\} : \hat{\mu}_i \geq \hat{\mu}_1] \\ &\leq \sum_{i=2}^n \Pr[\hat{\mu}_i \geq \hat{\mu}_1] \\ &\leq \sum_{i=2}^n \Pr\left[\hat{\mu}_i - \mu_i \geq \frac{\Delta_i}{2} \text{ or } \hat{\mu}_1 - \mu_1 \geq \frac{\Delta_i}{2}\right] \quad (\text{It can be understood from simple inspection}) \\ &\leq \sum_{i=2}^n \left(\Pr\left[\hat{\mu}_i - \mu_i \geq \frac{\Delta_i}{2}\right] + \Pr\left[\hat{\mu}_1 - \mu_1 \geq \frac{\Delta_i}{2}\right] \right) \\ &\leq \sum_{i=2}^n \left(\Pr\left[\hat{\mu}_i - \mu_i \geq \frac{\Delta_*}{2}\right] + \Pr\left[\hat{\mu}_1 - \mu_1 \geq \frac{\Delta_*}{2}\right] \right) \quad (\text{Because } \Delta_* \text{ is the minimum among } \Delta_i\text{s}) \\ &\leq 2(n-1) \exp\left(-2k \frac{\Delta_*^2}{4}\right) \\ &\leq 2(n-1) \exp\left(-2 \frac{4 \log(nT)}{\Delta_*^2} \frac{\Delta_*^2}{4}\right) \quad (\text{From (17.3)}) \\ &= 2(n-1) \frac{1}{n^2 T^2} \leq \frac{1}{T}. \end{aligned} \quad (17.6)$$

Therefore,

$$\mathbb{E}[N_i^{T+1}] \leq \frac{4 \log(nT)}{\Delta_*^2} + 1. \quad (17.7)$$

Finally,

$$\mathbb{E}[\text{regret}_T] \leq \sum_{i=2}^n \Delta_i \mathbb{E}[N_i^{T+1}]. \quad (17.8)$$

Hence,

$$\mathbb{E}[\text{regret}_T] \leq \sum_{i=2}^n \left[\frac{4\Delta_i \log(nT)}{\Delta_i^2} + O(\Delta_i) \right]. \quad (17.9)$$

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17.3 Upper Confidence Bound (UCB)

Consider the previous definition for $N_i^t = \sum_{s=1}^{t-1} \mathbb{1}\{i_s = i\}$, and the estimate of the mean for arm i up till the t^{th} round can be given by

$$\hat{\mu}_i = \sum_{s=1}^{t-1} \frac{X_i^s \mathbb{1}\{i_s = i\}}{N_i^t}. \quad (17.10)$$

The UCB algorithm chooses an arm at round t according to the relation

$$i_t = \arg \max_{i \in \{1, \dots, n\}} UCB_i^t, \quad (17.11)$$

where

$$UCB_i^t = \hat{\mu}_i + \sqrt{\frac{2 \log(1/\delta)}{N_i^t}}. \quad (17.12)$$

Here $\sqrt{\frac{2 \log(1/\delta)}{N_i^t}}$ is called the exploration bonus (optimism term). The term essentially increases the mean for an arm i if it has not been explored much, thus incentivizing exploration.

Theorem 17.2 For $\delta = 1/t^2$, the bound on the expected regret for the UCB algorithm over T rounds is given by

$$\mathbb{E}[\text{regret}_T] \leq \sum_{i=2}^n \left[\frac{16 \log T}{\Delta_i} + O\left(\sum_{i=2}^n \Delta_i\right) \right]. \quad (17.13)$$

The proof for the above theorem requires showing that bad arms are not chosen that often. Essentially, it has to be shown that $N_i^t > k_i$, where $k_i = \left\lceil \frac{8 \log(1/\delta)}{\Delta_i^2} \right\rceil$. Now, let $\mu_i^{\hat{k}_i} = \hat{\mu}_i^t$ when $N_i^t = k_i$ i.e., sample arm i enough times such that $\hat{\mu}_i^k$ is an empirical of k samples. We can define a good scenario as

$$G_i = \{\mu_1 < UCB_1^t, \forall t = 1, \dots, T\} \cap \left\{ \mu_i^{\hat{k}_i} + \sqrt{\frac{2 \log(1/\delta)}{k_i}} < \mu_1 \right\}. \quad (17.14)$$

Lemma 17.3 If G_i is true, then it is guaranteed that $N_i^{T+1} < k_i$.